#### The regression equation is

P/E = 6.70 + 0.183 Profit + 0.213 Growth + 0.84 Oil + 3.82 Drug

Predictor	Coef	StDev	Т	P
Constant	6.704	2.178	3.08	0.008
Profit	0.1827	0.2092	0.87	0.397
Growth	0.2128	0.1311	1.62	0.127
Oil	0.835	1.509	0.55	0.589
Drug	3.819	1.779	2.15	0.050

S = 2.309 R-Sq = 69.5% R-Sq(adj) = 60.8%

Analysis of Variance

Source Regressic Residual Total	on Error	DF 4 14 18	SS 169.887 74.650 244.537	MS 42.472 5.332	F 7.97	P 0.001
Source	DF	Se	a SS			
Profit	1	131	.398			
Growth	1	13	3.194			
Oil	1	C	0.724			
Drug	1	24	1.570			

Only the DRUG coefficient is significant at  $\alpha = 0.05$ .

(c) If we choose the drug/healthcare industry as the baseline, then we would instead have an indicator COMPUTER. There is no need to refit the model. The new coefficients, denoted with a subscript of N, depend on the previously fitted coefficients, denoted with a subscript of O, as below:

 $Constant_N = Constant_O + Drug_O = 6.704 + 3.819 = 10.523.$ 

Since now the constant term represents the baseline industry, DRUG, the PROFIT and GROWTH coefficients will not change.

 $OIL_N = OIL_O - DRUG_O = 0.835 - 3.819 = -2.984.$ 

 $COMPUTER_N = COMPUTER_O - DRUG_O = 0 - 3.819 = -3.819.$ 

Therefore the new model would be

 $\hat{y} = 10.523 + 0.183 \; \mathrm{PROFIT} \; + 0.213 \; \mathrm{GROWTH} \; - 2.984 \; \mathrm{OIL} \; - 3.819 \; \mathrm{COMPUTER}$  .

11.40 (a) The partial correlation coefficients are

$$r_{yx_1|x_2} = \sqrt{\frac{\text{SSE}(x_2) - \text{SSE}(x_1, x_2)}{\text{SSE}(x_2)}} = \sqrt{\frac{12.606 - 5.988}{12.606}} = 0.725,$$

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$$r_{yx_2|x_1} = \sqrt{\frac{\text{SSE}(x_1) - \text{SSE}(x_1, x_2)}{\text{SSE}(x_1)}} = \sqrt{\frac{13.519 - 5.988}{13.519}} = 0.746.$$

(b) The F statistics for testing the significance of these partial correlation coefficients are

$$F_{1} = \frac{\text{SSE}(x_{2}) - \text{SSE}(x_{1}, x_{2})}{\text{SSE}(x_{2})/(n-3)} = \frac{12.606 - 5.988}{5.988/(40-3)} = 40.90,$$
  
$$F_{2} = \frac{\text{SSE}(x_{1}) - \text{SSE}(x_{1}, x_{2})}{\text{SSE}(x_{1}, x_{2})/(n-3)} = \frac{13.519 - 5.988}{5.988/(40-3)} = 46.54.$$

Then

$$t_1 = \sqrt{40.90} = 6.395,$$
  
$$t_2 = \sqrt{46.54} = 6.822.$$

These the match the t statistics obtained in Exercise 11.2.

11.41 (a)  $r_{yx_1} = 0.378$ ,  $r_{yx_2} = -0.093$ , and  $r_{yx_3} = 0.003$ .  $x_1$  would enter first. (b) The partial correlation coefficients are

$$r_{yx_2|x_1} = \sqrt{\frac{\text{SSE}(x_1) - \text{SSE}(x_1, x_2)}{\text{SSE}(x_1)}} = \sqrt{\frac{16198 - 13322}{16198}} = 0.421,$$
  
$$r_{yx_3|x_1} = \sqrt{\frac{\text{SSE}(x_1) - \text{SSE}(x_1, x_3)}{\text{SSE}(x_1)}} = \sqrt{\frac{16198 - 15258}{16198}} = 0.241.$$

(c) The F statistics for testing the significance of these partial correlation coefficients are

$$F_{2} = \frac{\text{SSE}(x_{1}) - \text{SSE}(x_{1}, x_{2})}{\text{SSE}(x_{1}, x_{2})/(n-3)} = \frac{16198 - 13322}{13322/(38-3)} = 7.556,$$
  

$$F_{3} = \frac{\text{SSE}(x_{1}) - \text{SSE}(x_{1}, x_{3})}{\text{SSE}(x_{1}, x_{3})/(n-3)} = \frac{16198 - 15258}{15258/(38-3)} = 2.156.$$

Height  $(x_2)$  is the better predictor given that brain size is included in the model.

11.42 (a)  $r_{\log y, \log x_1} = -0.761$ ,  $r_{\log y, \log x_2} = -0.549$ , and  $r_{\log y, x_3} = -0.644$ .  $\log x_1$  would enter first.

(b) The partial correlation coefficients are

$$r_{\log y, \log x_2 | \log x_1} = \sqrt{\frac{\text{SSE}(\log x_1) - \text{SSE}(\log x_1, \log x_2)}{\text{SSE}(\log x_1)}} = \sqrt{\frac{6.112 - 5.815}{6.112}} = 0.220,$$
  
$$r_{\log y, x_3 | \log x_1} = \sqrt{\frac{\text{SSE}(\log x_1) - \text{SSE}(\log x_1, x_3)}{\text{SSE}(\log x_1)}} = \sqrt{\frac{6.112 - 6.053}{6.112}} = 0.098.$$

(c) The F statistics for testing the significance of these partial correlation coefficients are

$$F_{\log x_2} = \frac{\text{SSE}(\log x_1) - \text{SSE}(\log x_1, \log x_2)}{\text{SSE}(\log x_1, \log x_2)/(n-3)} = \frac{6.112 - 5.815}{5.815/(38-3)} = 1.788,$$
  
$$F_{x_3} = \frac{\text{SSE}(\log x_1) - \text{SSE}(\log x_1, x_3)}{\text{SSE}(\log x_1, x_3)/(n-3)} = \frac{6.112 - 6.053}{6.053/(38-3)} = 0.341.$$

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Log(calcium) is the better predictor, given that log(Alkalinity) is included in the model.

11.43

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$$F_{p} = \frac{(\text{SSE}_{p-1} - \text{SSE}_{p})/1}{\text{SSE}_{p}/[n - (p+1)]}$$

$$= \frac{(r_{yx_{p}|x_{1},...,x_{p-1}})(\text{SSE}_{p-1})}{\text{SSE}_{p}/[n - (k+1)]}$$

$$= \frac{r_{yx_{p}|x_{1},...,x_{p-1}}^{2}[n - (k+1)]}{\text{SSE}_{p}/\text{SSE}_{p-1}}$$

$$= \frac{r_{yx_{p}|x_{1},...,x_{p-1}}^{2}[n - (k+1)]}{1 - r_{yx_{p}|x_{1},...,x_{p-1}}^{2}}.$$

As  $r_{yx_p|x_1,...,x_{p-1}}^2$  increases, the numerator increases and the denominator decreases, so that  $F_p$  is an increasing function in  $r_{yx_p|x_1,...,x_{p-1}}^2$ .

**14 (a)** The stepwise regression output is shown below:

Stepwise Regression

F-to-Enter	r: 2	.00 F-	•to-Remove:	2.00	
Response :	is P/E	on 4	predictors,	with N =	19
Step	1	2			
Constant	10.309	8.319			
Drug	5.6	4.7			
T-Value	5.02	4.19			
Growth		0.23			
T-Value		1.97			
S	2.41	2.23			
R-Sq	59.73	67.60			

Drug enters at step 1, and Growth enters at step 2. The final model is  $\hat{y} = 8.319 + 4.7$  Drug + 0.23 Growth. To determine if these factors are significant at  $\alpha = 0.05$ , compare the t statistics to  $t_{19-3,0.025} = 2.120$ . Only Drug is significant.

(b) The best subsets regression output is shown below:

Best Subsets Regression

Response is P/E Ratio

Ρ	G	
r	r	
0	0	D

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					fwOr
		Adj.			itiu
Vars	R-Sq	R-Sq	C-p	S	thlg
1	59.7	57.4	3.5	2.4067	X
1	53.7	51.0	6.2	2.5798	X
1	32.1	28.1	16.1	3.1247	X
1	17.3	12.4	22.9	3.4500	X
2	67.6	63.5	1.9	2.2254	X X
2	63.5	59.0	3.7	2.3611	X X
2	59.8	54.8	5.4	2.4788 ·	XX
2	59.1	54.0	5.7	2.4993	XX
2	53.8	48.0	8.2	2.6582	X X
3	68.8	62.6	3.3	2.2551	XX X
3	67.8	61.4	3.8	2.2908	XXX
3	63.7	56.5	5.6	2.4316	X X X
3	59.4	51.3	7.6	2.5719	XXX
4	69.5	60.8	5.0	2.3091	XXXX
			-		

According to the  $C_p$  criterion, the best subset includes the Growth and Drug factors, with a  $C_p$  of 1.9. This is identical to the model found in part (a).

11.45 The best subsets regression output is given below:

Best Subsets Regression

Response is log(y)

					11
					0 0
					gg
					((
					хх
		Adj.			12 x
Vars	R-Sq	R-Sq	C-p	S	))3
4	<b>57 Q</b>	56.8	2.4	0.41203	X
1	A1 A	39.8	16.7	0.48625	X
1	30.2	28.2	26.4	0.53099	X
2	60.0	57.7	2.6	0.40759	XX
2	58.3	56.0	4.0	0.41588	X X
2	43.2	40.0	17.1	0.48547	XX
3	60.7	57.2	4.0	0.40988	XXX

Using the  $C_p$  criterion, the model with only  $\log x_1$  has the lowest  $C_p$ . This is the same model that was selected in Exercise 11.19. The fitted model is  $\log \hat{y} = 7.21 - 0.398 \log x_1$ , where  $\hat{\beta}_{\log x_1}$  was highly significant (*P*-value  $\approx 0.000$ ).

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			D df	MSE-	Adj. $r_2^2$	$C_p$
Variables in Model	SSE,	7	BITUR C.L.	morp		20
None	950	0	19	50	U	20
<b>T</b> 3	720	1	18	40	0.2	12.8
-1 	630	1	18	35	0.3	9.2
<i>L</i> <sub>2</sub>	540		18	30	0.4	5.6
#3	595	2	17	35	0.3	9.8
<i>x</i> <sub>1</sub> , <i>x</i> <sub>2</sub>	425	2	17	25	0.5	3
<i>w</i> <sub>1</sub> , <i>w</i> <sub>3</sub>	510	2	17	30	0.4	6.4
	400		16	25	0.5	4
<i>w</i> <sub>1</sub> , <i>w</i> <sub>2</sub> , <i>w</i> <sub>3</sub>						

(b) Subsets  $(x_1, x_3)$  and  $(x_1, x_2, x_3)$  have the maximum adjusted  $r_p^2$ . Subset  $(x_1, x_3)$  has the minimum  $C_p$ . Choose  $(x_1, x_3)$  since it has less variables and the minimum  $C_p$ .

(c)  $x_3$  gives the biggest reduction in  $SSE_p$  (no need to calculate partial F's for  $x_1, x_2, x_3$ ). So it will be the first variable to enter the model. The F to enter for  $x_3$  is

$$F_3 = \frac{950 - 540}{540/18} = 13.67.$$

Since  $F_3 > F_{IN} = 4.0$ ,  $x_3$  will enter the model.

(d)  $(x_1, x_3)$  gives the biggest reduction in  $SSE_p(x_3)$  (no need to calculate partial F for  $x_2$ . The F to enter for  $x_1$  is

$$F_{1|3} = \frac{SSE(x_3) - SSE(x_1, x_3)}{SSE(x_1, x_3)/(20 - 3)} = \frac{540 - 425}{425/17} = 4.6.$$

Since  $F_{1|3} > F_{IN}$ ,  $x_1$  will enter the model next. Its partial correlation is

$$r_{yx_1|x_3}^2 = \frac{SSE(x_3) - SSE(x_1, x_3)}{SSE(x_3)} = \frac{540 - 425}{540} = 0.213.$$

(e) The F to remove for  $x_3$  is

$$F_{3|1} = \frac{SSE(x_1) - SSE(x_1, x_3)}{SSE(x_1, x_3)/(20 - 3)} = \frac{720 - 425}{425/17} = 11.8.$$

Since  $F_{3|1} > F_{OUT} = 4.0$ ,  $x_3$  will not be removed.

(f) The F to enter for  $x_2$  is

$$F_{2|1,3} = \frac{SSE(x_1, x_3) - SSE(x_1, x_2, x_3)}{SSE(x_1, x_2, x_3)/(20 - 4)} = \frac{425 - 500}{400/16} = 1.0.$$

Since  $F_{2|1,3} < F_{IN} = 4.0$ ,  $x_2$  will not enter, and the full model will not be chosen.

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## Chapter 12 Solutions

 $\boldsymbol{s}$ 

Solutions to Section 12.1

12.1 (a) Sugar:

$$^{2} = \frac{(2.138)^{2} + (1.985)^{2} + (1.865)^{2}}{3} = 3.996,$$

so that  $s = \sqrt{3.996} = 1.999$  with  $3 \times 19 = 57$  d.f. Using the critical value  $t_{57,0.025} \approx 2.000$ , the 95% CI's are :

Shelf 1 : 
$$4.80 \pm (2.000) \times \frac{1.999}{\sqrt{20}} = [3.906, 5.694]$$
  
Shelf 2 :  $9.85 \pm (2.000) \times \frac{1.999}{\sqrt{20}} = [8.956, 10.744]$   
Shelf 3 :  $6.10 \pm (2.000) \times \frac{1.999}{\sqrt{20}} = [5.206, 6.994].$ 

Fiber:

$$s^{2} = \frac{(1.166)^{2} + (1.162)^{2} + (1.277)^{2}}{3} = 1.447,$$

so that  $s = \sqrt{1.447} = 1.203$  with  $3 \times 19 = 57$  d.f. Using the critical value  $t_{57,0.025} \approx 2.000$ , the 95% CI's are :

Shelf 1 : 
$$1.68 \pm (2.000) \times \frac{1.203}{\sqrt{20}} = [1.142, 2.218]$$
  
Shelf 2 :  $0.95 \pm (2.000) \times \frac{1.203}{\sqrt{20}} = [0.412, 1.488]$   
Shelf 3 :  $2.17 \pm (2.000) \times \frac{1.203}{\sqrt{20}} = [1.632, 2.708].$ 

Shelf 2 cereals are higher in sugar content than shelves 1 and 3, since the CI for shelf 2 is above those of shelves 1 and 3. Similarly, the shelf 2 fiber content CI is below that **a** shelf 3. So in general, shelf 2 cereals are higher in sugar and lower in fiber.

(b) Sugar:

$$SSE = 57 \times 3.996 = 227.80,$$
  

$$SSA = n \sum \bar{y}_i^2 - N \bar{y}^2$$
  

$$= 20[(4.80)^2 + (9.85)^2 + (6.10)^2] - 60 \times (6.92)^2$$
  

$$= 275.03.$$

Then the ANOVA table is below:

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Since  $F > f_{2,57,0.05} = 3.15$ , there do appear to be significant differences among the shelves in terms of sugar content.

Fiber:

$$SSE = 57 \times 1.447 = 82.47,$$
  

$$SSA = n \sum \bar{y}_i^2 - N \bar{y}^2$$
  

$$= 20[(1.68)^2 + (0.95)^2 + (2.17)^2] - 60 \times (1.60)^2$$
  

$$= 15.08.$$

Then the ANOVA table is below:

Analysis of Variance							
Source	SS	d.f.	MS	F			
Shelves	15.08	2	7.54	5.21			
Error	82.47	57	1.447				
Total	97.55	59					

Since  $F > f_{2,57,0.05} = 3.15$ , there do appear to be significant differences among the shelves in terms of fiber content, as well as sugar content.

(c) The grocery store strategy is to place high sugar/low fiber cereals at the eye height of school children where they can easily see them.

**12.2** (a)



The boxplot indicates that the number of finger taps increases with higher doses of caffeine.

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(b)

Analysis of Variance							
Source SS d.f. MS F							
Dose	61.400	2	30.700	6.181			
Error	134.100	27	4.967				
Total	195.500	29					

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Since  $F > f_{2,27,0.05} = 3.35$ , there do appear to be significant differences in the numbers of finger taps for different doses of caffeine.



From the plot of the residuals against the predicted values, the constant variance assumption appears satisfied. From the normal plot of the residuals, the residuals appear to follow the normal distribution.

12.3 (a)



The boxplot indicates that the control is higher than the two treatments in the average . number of eggs laid.

(b)

2,519

(c)



The boxplot indicates that HBSC has the highest average hemoglobin level, followed by HBS, and then HBSS.

Analysis of Variance							
Source SS d.f. MS F							
Disease type	99.889	2	49.945	49.999			
Error	37.959	38	0.999				
Total	137.848	40					

Since  $F > f_{2,38,0.05} = 3.23$ , there do appear to be significant differences in the hemoglobin **levels** between patients with different types of sickle cell disease.



From the plot of the residuals against the predicted values, the constant variance assumption appears satisfied. From the normal plot of the residuals, the residuals appear to follow the normal distribution.

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From the plot of the residuals against the predicted values, the constant variance assumption appears satisfied. From the normal plot of the residuals, the residuals appear to follow the normal distribution.



The side-by-side boxplots of the pre-treatment weights overlap quite a bit, and don't indicate large differences among the groups.

(b)

**I** (a)

Analysis of Variance							
Source SS d.f. MS F							
Treatment	31.2	2	15.6	0.50			
Error	1503.5	48	31.0				
Total	1534.7	50					

Since  $F < f_{2,48,0.05} = 3.19$ , the pre-treatment weights are not significantly different among the different treatment groups.

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(c)



Here the differences among the treatment groups are more pronounced. The family group appears to be different than the other two, but it is hard to tell if the differences are significant.

(d)

			the second se	
Ar	alvsis of	Varia	nce	
Cource	SS	d.f.	MS	F
Source	479.3	2	239.7	3.85
Treatment	2990.6	48	62.3	
Error	2460.0	50		
lotal	0409.5			

Since  $F > f_{2,48,0.05} = 3.19$ , the weight differences are significantly different among the different treatment groups.

# Solutions to Section 12.2

12.8 Sugar: The number of comparisons is

$$\binom{a}{2} = \binom{3}{2} = 3,$$

Then

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$$t_{57,\frac{0.05}{2\times3}} = t_{57,0.0083} = 2.468,$$

and the Bonferroni critical value is

$$t_{57,0.0083}s\sqrt{\frac{2}{n}} = 2.468 \times 1.999\sqrt{\frac{2}{20}} = 1.56.$$

For the Tukey method,

 $q_{3,57,0.05} \approx 3.40,$ 

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and the Tukey critical value is

$$q_{3,57,0.05}s\sqrt{\frac{1}{n}} = 3.40 \times 1.999\sqrt{\frac{1}{20}} = 1.52.$$

Since  $t_{13} > 2.051$ , conclude that sites 1 and 3 are significantly different. Similarly, since  $t_{23} > 2.051$ , conclude that sites 2 and 3 are significantly different. So the conclusion is the same as with the Tukey method, namely that all three sites are significantly different from one another.

12.10 The number of comparisons is

$$\binom{a}{2} = \binom{3}{2} = 3,$$

Then

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$$t_{38,\frac{0.01}{2\times3}} = t_{38,0.0017} = 3.136,$$

and, since  $s = \sqrt{0.999} = 1$ , the form of the Bonferroni confidence interval is

$$\bar{y}_i - \bar{y}_j \pm (3.136)(1)\sqrt{\frac{1}{n_i} + \frac{1}{n_j}}.$$

For the Tukey method,

$$\frac{q_{3,38,0.01}}{\sqrt{2}} \approx \frac{4.39}{\sqrt{2}} = 3.104,$$

and the form of the Tukey confidence interval is

$$\bar{y}_i - \bar{y}_j \pm (3.104)(1)\sqrt{\frac{1}{n_i} + \frac{1}{n_j}}.$$

For the LSD method,  $t_{38,0.01/2} = 2.712$ , and the form of the LSD confidence interval is

$$\bar{y}_i - \bar{y}_j \pm (2.712)(1) \sqrt{\frac{1}{n_i} + \frac{1}{n_j}}.$$

The 99% confidence intervals are summarized in the table below:

		Bonferroni		Tukey		LSD	
Comparison	$ ar{y}_i - ar{y}_j $	Lower	Upper	Lower	Upper	Lower	Upper
HBSC(3) vs. $HBS(2)$	1.670	0.392	2.948	0.406	2.934	0.564	2.776
HBS(2) vs. $HBSS(1)$	1.918	0.656	3.179	0.669	3.166	0.825	3.010
HBSC(3) vs. $HBSS(1)$	3.588	2.462	4.713	2.475	4.700	2.614	4.561

Since all of these intervals are entirely above 0, all of the types of disease have significantly different hemoglobin levels from one another. Note that the Bonferroni method has the widest intervals, followed by the Tukey. The LSD intervals are narrowest because there is no adjustment for multiplicity.

12.11 The ANOVA output, including Tukey 90% confidence intervals, is given below:

One-way Analysis of Variance

Analysis	of Vari	ance for	Height		
Source	DF	SS	MS	F	Р
Part	3	81.11	27.04	3.95	0.010

705.67 6.85 103 Error 786.79 Total 106 Individual 95% CIs For Mean Based on Pooled StDev N Mean StDev \_\_\_\_\_ \_\_\_\_\_ Level (-----) 39 70.718 2.361 Bass1 71.385 2.729 (-----) 26 Bass2 (-----) 68.905 3.330 21 Tenor1 (-----) 69.905 2.071 21 Tenor2 \_\_\_\_\_ 69.0 70.5 72.0 Pooled StDev = 2.617 Tukey's pairwise comparisons Family error rate = 0.100 Individual error rate = 0.0224 Critical value = 3.28 Intervals for (column level mean) - (row level mean) Bass2 Tenor1 Bass1 -2.204Bass2 0.870 0.699 0.170 Tenor1 4.261 3.456 -2.873 -0.301 -0.830 Tenor2 3.261 0.873 2.456

The only significant difference is that Tenor 1 men are shorter than both Bass 1 and Bass 2 men on average, since those confidence intervals are entirely above 0.

12.12 Since we suspect that caffeine will increase the rate of finger taps, the one-sided Dunnett critical value is

$$t_{a-1,\nu,\alpha} = t_{2,27,0.10} \approx 1.625.$$

Then the Dunnett lower confidence bounds, using  $s = \sqrt{4.97} = 2.23$ , are

$$\begin{array}{rcl} \mu_i - \mu_1 & \geq & \bar{y}_i - \bar{y}_1 - t_{2,27,0.10} s \sqrt{2/n} \\ \mu_2 - \mu_1 & \geq & 246.40 - 244.80 - (1.625)(2.23) \sqrt{2/10} = -0.021 \\ \mu_3 - \mu_1 & \geq & 248.30 - 244.80 - (1.625)(2.23) \sqrt{2/10} = 1.879. \end{array}$$

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Only the 200 mg dose is significantly higher than the control, since the confidence interval is entirely above 0.

12.13

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Player	li	u <sub>i</sub>
Jordan	$ 3.54 - 5.82 - 1.668 ^{-} = -3.948$	$ 3.54 - 5.82 + 1.668 ^+ = 0$
Rodman	$ 2.55 - 5.82 - 1.668 ^{-} = -4.938$	$ 2.55 - 5.82 + 1.668 ^+ = 0$
Kukoc	$ 5.82 - 3.54 - 1.668 ^{-} = 0$	$ 5.82 - 3.54 + 1.668 ^+ = 3.948$
Longley	$ 3.09 - 5.82 - 1.668 ^{-} = -4.398$	$ 3.09 - 5.82 + 1.668 ^+ = 0$
Harper	$ 2.82 - 5.82 - 1.668 ^{-} = -4.668$	$ 2.82 - 5.82 + 1.668 ^+ = 0$

Since Kukoc is the only player with  $l_i = 0$ , he does appear to have the most assists.

#### Solutions to Section 12.3

12.18 (a) Since

then

$$SSE = MSE \times (N - a) = 10.533 \times (120 - 10) = 1158.63$$
  

$$SSA = n \sum \bar{y}_i^2 - N \bar{y}^2$$
  

$$= 12[(36.76)^2 + ... + (32.98)^2] - 120 \times 34.454$$
  

$$= 279.932, \text{ and}$$
  

$$SST = SSE + SSA = 1158.63 + 279.932 = 1438.562.$$

 $MSE = \frac{s_1^2 + \ldots + s_{10}^2}{10} = 10.533,$ 

Then the ANOVA table is given below:

Analysis of Variance					
Source	SS	d.f.	MS	F	
Batch	279.932	9	31.104	2.953	
Error	1158.630	110	10.533		
Total	1438.562	119			

(b) The variance components estimates are

$$\hat{\sigma}_{\text{Error}}^2 = MSE = 10.533 \text{ and}$$
  
 $\hat{\sigma}_{\text{Batch}}^2 = \frac{MSA - MSE}{n}$   
 $= \frac{31.104 - 10.533}{12} = 1.714.$ 

Batch to batch variation accounts for about

$$\frac{1.714}{1.714 + 10.533} = 14\%$$

of the total error variation.

12.19 (a)

:

$$\operatorname{Var}(\bar{Y}) = \operatorname{Var}\left(\frac{1}{n}\sum_{j=1}^{n}Y_{ij}\right)$$

$$= \operatorname{Var}\left(\frac{1}{n}\sum_{j=1}^{n}(\mu + \tau_{i} + \epsilon_{ij})\right)$$
$$= \operatorname{Var}\left(\mu + \tau_{i} + \frac{1}{n}\sum_{j=1}^{n}\epsilon_{ij}\right)$$
$$= \operatorname{Var}(\tau_{i}) + \frac{1}{n^{2}}\sum_{j=1}^{n}\operatorname{Var}(\epsilon_{ij})$$
$$= \sigma_{\operatorname{Batch}}^{2} + \frac{\sigma_{\operatorname{Error}}^{2}}{n}.$$

(b) Using

$$\hat{\sigma}_Y = \sqrt{\hat{\sigma}_{\text{Batch}}^2 + \frac{\hat{\sigma}_{\text{Error}}^2}{12}} = \sqrt{1.714 + \frac{10.533}{12}} = \sqrt{2.592} = 1.610,$$

the three sigma control limits are

 $34.454 \pm 3(1.610) = [29.624, 39.284].$ 

Since all the batch means fall within these limits, the process is under control.

12.20 From the ANOVA table given in Exercise 12.16, the variance components estimates are

$$\hat{\sigma}_{\text{Error}}^2 = MSE = 26.53 \text{ and}$$
  
 $\hat{\sigma}_{\text{Cable}}^2 = \frac{MSA - MSE}{n}$   
 $= \frac{240.54 - 26.53}{12} = 17.834$ 

Cable to cable variation accounts for about

$$\frac{17.834}{17.834 + 26.53} = 40.2\%$$

of the total error variation.

## Solutions to Section 12.4

12.21 (a) Since the overall mean is 86, the effect estimates are

Quantity	Estimate	Quantity	Estimate
$ au_A$	84 - 86 = -2	$\beta_1$	92 - 86 = 6
$ au_B$	85 - 86 = -1	$\beta_2$	83 - 86 = -3
$ au_C$	89 - 86 = 3	$\beta_3$	85 - 86 = -1
$ au_D$	86 - 86 = 0	$\beta_4$	88 - 86 = 2
		$eta_5$	82 - 86 = -4

(b)

·	Analys	is of <b>\</b>	<b>Variance</b>	
Source	SS	d.f.	MS	$\mathbf{F}$
Blend	264	4	66.000	3.504
Method	. 70	3	23.333	1.239
Error	226	12	18.833	
Total	560	19		· · · · · · · · · · · · · · · · · · ·

For method,  $F < f_{3,12,0.05} = 3.49$ , so there do not appear to be any significant differences between methods. For blend,  $F > f_{4,12,0.05} = 3.26$ , so there do appear to be significant differences between blends.



There is possibly a decreasing variance with increasing yield, but this is unclear because there are also less observations at higher yields, so the distribution is less likely to be filled out. From the normal plot of the residuals, the residuals appear to follow the normal distribution.

12.22 (a) The ANOVA output is given below:

General Linear Model

Factor Player Stadium	Typ fixe fixe	e Levels Va ed 6 Be ed 3 Ho	lues cker Cole me Othe	Cordova r Outdoo	Mack rs	Munoz	Puckett
Analysis	of Va	riance for	Zone Rat, u	sing Adjust	ed SS fo	or Tests	
Source	DF	Seq SS	Adj SS	Adj MS	F	Р	
Player	5	0.0148578	0.0148578	0.0029716	23.79	0.000	
Stadium	2	0.0016591	0.0016591	0.0008296	6.64	0.015	
Error	10	0.0012489	0.0012489	0.0001249			
Total	17	0.0177658					

-

Least Squares Means for Zone Rat

Stadium	Mean	StDev	
Home	0.8007	0.004562	•-
Other	0.8020	0.004562	

Outdoors 0.8217 0.004562

Tukey 95.0% Simultaneous Confidence Intervals Response Variable Zone Rat All Pairwise Comparisons among Levels of Stadium

Stadium = Home subtracted from:



-0.015

0.000

------

0.030

0.015

Since the *P*-value for Stadium is  $0.015 < \alpha = 0.05$ , conclude that there are significant differences between the stadiums.

- (b) The Tukey confidence intervals are given in the computer output from (a). From whether these confidence intervals contain 0 or not, we can see that Outdoor stadiums are significantly different from both Home Domes and Other Domes, but Home Domes are not significantly different from Other Domes.
- 23 The estimated contrast is

$$c = \frac{1}{2}(\bar{y}_1 + \bar{y}_2) - \bar{y}_3 = \frac{1}{2}(0.8007 + 0.802) - 0.8217 = -0.020,$$

with a standard error of

s.e.
$$(c) = \sqrt{s^2 \sum_{i} \left(\frac{c_i^2}{n_i}\right)} = \sqrt{0.000125 \left(\frac{(0.5)^2 + (0.5)^2 + (-1)^2}{6}\right)} = 0.0056.$$

Then the test statistic is

$$t = \frac{-0.020}{0.0056} = -3.578.$$

Since  $|t| > t_{10,0.025} = 2.228$ , conclude that there is a significant difference between domed stadiums and outdoor stadiums.

24 (a) The ANOVA output is given below:

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Analysis of Variance						
Source	SS	d.f.	MS	$\mathbf{F}$		
Park	0.170	1	0.170	14.130		
Year	1.064	7	0.152	12.627		
Error	0.084	7	0.012			
Total	1.319	15				

Since  $F_{\text{Park}} > f_{1,7,0.05} = 5.59$ , there do appear to be significant differences between new and unchanged parks in their HR/G ratios.

- 12.28 (a) If batches were ignored as a blocking factor, the SS and d.f. for batches would both be included in the error SS and d.f.
  - (b) The new ANOVA table would be

Analysis of Variance					
Source	SS	d.f.	MS	F	
Position	40.396	7	5.771	3.614	
Error	25.549	16	1.597		
Total	65.945	23			

Since  $F < f_{7,16,0.01} = 4.03$ , the positions are not significantly different from one another, when the batches are ignored as a blocking factor.

(c) A nonsignificant result is obtained in (b) because the variance associated with blocks was no longer removed and was included in the error variance. This raised the denominator, reduced the F statistic, and made it more difficult to detect differences among the positions. However, this will not always happen. If the  $MS_{Batches}$  is smaller than MSE, including the batch to batch variation in the error term will decrease the overall MSE and have the opposite effect. Also, including the batch d.f. in the error d.f. will make the critical F value smaller, and make it easier to reject  $H_0$ .

### 12.29 Since there are 6 blocks, n = 6 observations per variety. Also, since

$$s = \sqrt{79.64} = 8.924$$

then

 $d = t_{7-1,30,0.10} s \sqrt{\frac{2}{n}} = 2.046 \times 8.924 \times \sqrt{\frac{2}{6}} = 10.542.$ 

Using

$$\begin{split} l_i &= |\bar{y}_i - \max_{j \neq i} \bar{y}_j - d|^-, \\ u_i &= |\bar{y}_i - \max_{i \neq i} \bar{y}_j + d|^+, \end{split}$$

Variety	li	u <sub>i</sub>
Α	$ 49.6 - 71.3 - 10.542 ^{-} = -32.242$	$ 49.6 - 71.3 + 10.542 ^+ = 0$
В	$ 71.2 - 71.3 - 10.542 ^{-} = -10.642$	$ 71.2 - 71.3 + 10.542 ^+ = 10.442$
С	$ 67.6 - 71.3 - 10.542 ^{-} = -14.242$	$ 67.6 - 71.3 + 10.542 ^+ = 6.842$
D	$ 61.5 - 71.3 - 10.542 ^{-} = -20.342$	$ 61.5 - 71.3 + 10.542 ^+ = 0.742$
E	$ 71.3 - 71.2 - 10.542 ^{-} = -10.442$	$ 71.3 - 71.2 + 10.542 ^+ = 10.642$
F	$ 58.1 - 71.3 - 10.542 ^{-} = -23.742$	$ 58.1 - 71.3 + 10.542 ^+ = 0$
G	$ 61.0 - 71.3 - 10.542 ^{-} = -20.842$	$ 61.0 - 71.3 + 10.542 ^+ = 0.242$

the results are summarized in the table below: