Wavelet approximations to Jacobians
and the inversion of complicated maps

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Abstract
Principal orthogonal decomposition can be used to solve two related problems: distinguishing elements from a
collection by making \(d\) measurements, and inverting a complicated map from a \(p\)-parameter configuration space to
a \(d\)-dimensional measurement space. In the case where \(d\) is more than 1000 or so, the classical \(O(d^3)\) singular value
decomposition algorithm becomes very costly, but it can be replaced with an approximate best-basis method that
has complexity \(O(d^2 \log d)\). This can be used to compute an approximate Jacobian for a complicated map from
\(\mathbb{R}^p \rightarrow \mathbb{R}^d\) in the case \(p \ll d\). This paper is a condensed version of an invited lecture given at Math-Chem-Comp

1 Introduction
Consider the problem of most efficiently distinguishing elements from a collection by making \(d\) measurements.
In general, we will need all \(d\) measured values to fully specify an element. However, it is possible to use less
information if some of the measurements are correlated. For example, if the objects are parameterized by a small
number \(p \ll d\) of parameters, then the \(d\) measurements should separate them in a redundant fashion. That is,
we can change basis locally in the \(d\)-dimensional measurement space to find just \(p\) combinations of measurements
which change with the \(p\) parameters. This idea works even if there are many parameters but only \(p\) of them are
relatively important.

Note the resemblance between the problem of distinguishing elements and the problem of inverting a complicated
map from \(\mathbb{R}^p\) to \(\mathbb{R}^d\). In the first problem, we must find a discrete object given its description in \(\mathbb{R}^d\). In
the second, we must find the parameters in \(\mathbb{R}^p\) from the description in \(\mathbb{R}^d\). These problems are identical if the
collection of objects is produced by evaluating the complicated map at discrete grid points in \(\mathbb{R}^p\).

The combinations of measurements which root out the underlying parameters are called principal (orthogonal)
components or factors; they have a precise meaning, and the well-known and well-behaved method of singular
value decomposition or SVD produces them with arbitrary accuracy. However, SVD has a complexity that is
asymptotically \(O(d^3)\), making it impractical for problems larger than \(d \approx 1000\). In this paper, we will describe the
classical principal factor algorithm, then give a lower-complexity approximate principal factor algorithm. Finally,
we will give two example applications of the approximate algorithm to the two mentioned problems.

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2 Notation and definitions

We need some definitions from probability theory.

2.1 Random vectors, means and variances

Let $X = \{X_n : n = 1, \ldots, N\} \subset \mathbb{R}^d$ be an ensemble of vectors. Write

$$E(X) \overset{\text{def}}{=} \frac{1}{N} \sum_{n=1}^{N} X_n$$

for the average vector in the ensemble, i.e., the expectation of $x$ over the set $X$.

Let $\sigma(X) \subset \mathbb{R}^d$ be the vector of the standard deviations of the coefficients of $X$. Namely,

$$\sigma(X) = \sqrt{E(X^2) - E(X)^2}$$

$$\sigma(X)(k) = \left( \frac{1}{N} \sum_{n=1}^{N} [X_n(k) - E(X)(k)]^2 \right)^{1/2}$$

We define the variance ellipsoid of an ensemble $X$ to be the ellipsoid centered at $E(X) \in \mathbb{R}^d$, with semiaxes $\sigma(X)(1), \sigma(X)(2), \ldots, \sigma(X)(d)$ aligned with the $d$ coordinate axes. Its volume is $\omega_d \times [\sigma(X)(1)] \times [\sigma(X)(2)] \times \cdots \times [\sigma(X)(d)]$, where $\omega_d$ is the surface area of the unit sphere in $\mathbb{R}^d$.

If the ensemble $X$ is fixed forever, then we can assume without loss of generality that $\frac{1}{N} \sum_{n=1}^{N} X_n(k) = 0$ for all $k = 1, 2, \ldots, d$, namely that $E(X) = 0$, because this can always be arranged by subtracting the average vector $E(X)$ from each of $X_1, X_2, \ldots, X_N$. It results in a simpler formula for $\sigma(X)$:

$$E(X) = 0 \Rightarrow \sigma(X)(k) = \left( \frac{1}{N} \sum_{n=1}^{N} X_n(k)^2 \right)^{1/2}$$

Then the variance ellipsoid is centered at 0.

We write $\text{Var}(X)$ for the total variance of the ensemble $X$. This is the sum of the squares of the coordinates in the variance vector $\sigma(X) \in \mathbb{R}^d$. In other words, $\text{Var}(X) = \|\sigma(X)\|^2 \overset{\text{def}}{=} \sum_{k=1}^{d} \sigma(X)(k)^2$, or

$$\text{Var}(X) = \sum_{k=1}^{d} \left[ \frac{1}{N} \sum_{n=1}^{N} X_n(k)^2 - \left( \frac{1}{N} \sum_{n=1}^{N} X_n(k) \right)^2 \right].$$

2.2 The Karhunen–Loève transform

The autocovariance matrix for the ensemble $X$ is defined by

$$M \overset{\text{def}}{=} E(\bar{X} \otimes \bar{X}); \quad M(i,j) = \frac{1}{N} \sum_{n=1}^{N} \bar{X}_n(i) \bar{X}_n(j).$$

Here we have taken $\bar{X}_n \overset{\text{def}}{=} X_n - E(X)$ to be the original vector with the average vector subtracted. Thus $E(\bar{X}) = 0$. The matrix coefficient $M(i,j)$ is the covariance of the $i^{\text{th}}$ and $j^{\text{th}}$ coordinate of the random vector $\bar{X}$,
using the ensemble as the sample space. The matrix $M$ is evidently symmetric. It is also positive (some would say positive semidefinite) since for every vector $Y \in \mathbb{R}^d$ we have

$$
\langle Y, MY \rangle = \sum_{i=1}^{d} \sum_{j=1}^{d} Y(i)M(i,j)Y(j)
$$

$$
= \frac{1}{N} \sum_{n=1}^{N} \sum_{i=1}^{d} \sum_{j=1}^{d} Y(i)\bar{X}_n(i)\bar{X}_n(j)Y(j)
$$

$$
= \frac{1}{N} \sum_{n=1}^{N} \langle Y, \bar{X}_n \rangle^2 \geq 0
$$

Therefore, we can find an orthonormal basis for $\mathbb{R}^d$ consisting of eigenvectors of the matrix $M$, and these eigenvectors will have positive real eigenvalues. Thus we are assured that the Karhunen–Loève basis exists for the ensemble $X$ of vectors; this is the orthonormal basis of eigenvectors of the autocovariance matrix $E(\bar{X} \otimes \bar{X})$.

The Karhunen–Loève basis eigenvectors are also called principal orthogonal components or principal factors, and computing them for a given ensemble $X$ is also called factor analysis. Since the autocovariance matrix for the Karhunen–Loève eigenvectors is diagonal, it follows that the coordinates of the vectors in the sample space $X$ with respect to the Karhunen–Loève basis are uncorrelated random variables. Let us denote these basis eigenvectors by $\{Y_n : n = 1, \ldots, N\}$, and let us denote by $K$ the $d \times d$ matrix whose columns are the vectors $Y_1, \ldots, Y_N$. The adjoint of $K$, or $K^*$, is the matrix which changes from the standard coordinates into Karhunen–Loève coordinates; this map is called the Karhunen–Loève transform.

Unfortunately, finding these eigenvectors requires diagonalizing a matrix of order $d$, which has complexity $O(d^3)$. In addition, even after already computing the Karhunen–Loève eigenvectors of an ensemble, updating the basis with some extra random vectors will cost an additional $O(d^3)$ operations since it requires another
diagonalization.

Such a high order of complexity imposes a ceiling on the size of problem we can do by this method. In many cases of interest, \( d \) is very large and \( X \) spans \( \mathbb{R}^d \), implying \( N \geq d \). Thus even to find the coefficients of the autocovariance matrix requires \( O(d^3) \) operations. At the present time, we are limited to \( d \leq 10^3 \) if we must use common desktop computing equipment, and \( d \leq 10^4 \) for the very most powerful computers.

So we shall take another perspective. We shall pose the problem of finding the Karhunen-Loève eigenvectors as an optimization over the set of orthogonal transformations of the original ensemble \( X \). The quantity to be maximized will be a transform coding gain, or the amount of compression we achieve by using another basis to represent the ensemble. Alternatively, we could simply introduce a distance measure on the set of orthonormal bases for \( \mathbb{R}^d \), treat the Karhunen-Loève basis as a distinguished point in this set and then try to find an efficiently-computable basis which is close to the Karhunen-Loève basis.

\section*{2.3 A metric on the orthogonal matrices}

Consider Figure 2, which schematically depicts all the orthonormal bases of \( \mathbb{R}^d \). These can be identified with orthogonal transformations of \( \mathbb{R}^d \). The points marked by “\( x \)” represent bases in some library of fast transforms. The point marked “\( o \)” represents the optimal, or Karhunen-Loève basis, for a given ensemble of vectors. The point marked “\( xx \)” represents the fast transform closest to the Karhunen-Loève bases.

Let \( U \) be an orthogonal \( d \times d \) matrix, and write \( Y = UX \) to signify that \( Y_n = UX_n \) for each \( n = 1, 2, \ldots, N \). Since \( U \) is linear, \( E(Y) = E(UX) = UE(X) \), which will be 0 if we started with \( E(X) = 0 \). Since \( U \) is orthogonal, it preserves sums of squares, so \( \text{Var}(X) = \text{Var}(Y) \). Using wavelet packets \cite{1} or adapted local trigonometric functions \cite{3}, it is possible to build a library of more than \( 2^d \) fast transforms \( U \) of \( \mathbb{R}^d \) to use for the “\( x \)” points. We will illustrate the construction using the simple Haar-Walsh wavelet packets in the example below. These transforms are arranged in a structure that permits us to search for the one closest to the “\( o \)” point in \( O(d \log d) \)

\begin{figure}[h]
\centering
\includegraphics[width=0.5\textwidth]{basis.png}
\caption{Orthonormal bases for \( \mathbb{R}^d \)}
\end{figure}

\( x \) = fast transform bases; \( o \) = Karhunen-Loève basis; \( xx \) = best fast basis.
operations. We will use a notion of closeness that is derived from the function minimized by the Karhunen–Loève transform.

As in [4], define the transform coding gain for an orthogonal matrix by the formula

$$G_{TC}(U) = \text{Var}(UX)/\exp H(UX), \quad \text{where} \quad H(X) = \frac{1}{d} \sum_{i=1}^{d} \log \sigma(X)(i).$$

(7)

From this formula we see that $G_{TC}(UX)$ is maximized when $H(UX)$ is minimized. The quantity $H$ has various interpretations. It is the entropy of the direct sum of $d$ independent Gaussian random variables with variances $\sigma(X)(i), i = 1, \ldots, d$. It is also equal to the logarithm of the volume of the variance ellipsoid (if we add $\log \omega_d$), so we see that minimizing $H(UX)$ or maximizing $G_{TC}(UX)$ is equivalent to minimizing the volume of the variance ellipsoid for the ensemble $UX$ over all orthogonal matrices $U$.

The Karhunen–Loève transform is therefore a global minimum for $H$, and we will say that the best approximation to the Karhunen–Loève transform from a library $\mathcal{U}$ of orthogonal matrices is the minimum of $H(UX)$ with the constraint $U \in \mathcal{U}$. We can define the approximate factor analysis algorithm to be the searches through a library of orthonormal bases for the one closest to the Karhunen–Loève basis. If the library of bases is organized to facilitate a fast search, we will dare to call the result a fast approximate Karhunen–Loève algorithm.

The “closeness” of a basis $U$ to the Karhunen–Loève basis $K$ can be measured by computing the transform coding gain of $U$ and subtracting that of $K$. This gives us a transform coding gain metric:

$$\delta_X(U, V) = |H(UX) - H(VX)|$$

Notice that we get a different metric for each ensemble $X$. This is a degenerate metric on the orthogonal group, since it gives a distance of 0 between bases which have the same transform coding gain for $X$. However, this technical point can be overcome by constructing a topological quotient space in which such bases are considered equivalent.

A minimum for $H(VX)$ is the Karhunen–Loève basis $V = K$, so minimizing $\delta_X(U, K)$ over fast transforms $U$ finds the closest fast transform for this ensemble in the transform coding sense.

3 The approximate KL transform

We will now describe a large library of rapidly-computable orthonormal bases, constructed by a recursive decomposition algorithm which takes advantage of the rapid growth in the number of subtrees of a binary tree. We also describe an efficient search algorithm which can be used to maximize transform coding gain (or minimize entropy) over all the bases which fall from the tree. The vectors which make up these bases are discrete approximations to wavelet packets, which are described in detail in several papers [1, 3, 2]. We will fix our attention on the Haar–Walsh wavelet packets. It is relatively easy to generalize this example to the other wavelet packet approximate Karhunen–Loève expansions.

Let $S$ and $D$ be two operators which, together with their respective adjoints $S^*$ and $D^*$, are defined on sequences by the following formulas:

$$Sx(n) = \frac{[x(2n) + x(2n + 1)]}{\sqrt{2}};$$

$$S^* x(n) = \begin{cases} \frac{1}{\sqrt{2}} x\left(\frac{n}{2}\right), & \text{if } n \text{ is even}, \\ \frac{1}{\sqrt{2}} x\left(\frac{n-1}{2}\right), & \text{if } n \text{ is odd}; \end{cases}$$

$$Dx(n) = \frac{[x(2n) - x(2n + 1)]}{\sqrt{2}};$$

$$D^* x(n) = \begin{cases} \frac{1}{\sqrt{2}} x\left(\frac{n}{2}\right), & \text{if } n \text{ is even}, \\ -\frac{1}{\sqrt{2}} x\left(\frac{n-1}{2}\right), & \text{if } n \text{ is odd}. \end{cases}$$
$S$ is the so-called low-pass filter and $D$ is the high-pass filter conjugate quadrature filter. They are also called Haar-Walsh filters, since they are used to produce Haar and Walsh bases. The factors $1/\sqrt{2}$ ensure that the operators conserve energy and variance. $S^*$ and $D^*$ are the adjoint filters, and it is easy to see that $SS^* = DD^* = I$, $SD^* = DS^* = 0$ and $S^*S + D^*D = I$. Those facts are used to establish some remarkable orthogonality relations among wavelet packets.

Figure 3: Decomposition into Haar–Walsh wavelet packets

Starting with a signal $x = \{x(n) : n = 0, 1, \ldots, d - 1\}$ of $d = M2^L$ samples, we can recursively apply $S$ and $D$ a total of $L$ times because the number of samples is divisible by 2 at least $L$ times. We arrange the resulting sequences in a binary tree as in Figure 3. Notice that if we take a “graph” of blocks from that tree, namely a collection with the property that each vertical line goes through exactly one block, then the numbers in the blocks are coefficients with respect to an orthonormal basis. Each graph gives a different orthonormal basis; for example, the shaded blocks depicted in Figure 4 contain the so-called “wavelet basis” coefficients.

We then sum the coefficients of each tree into two “accumulator” trees:

- the tree of means, which contains $\sum_{n=0}^{N-1} DSSD \cdots Dx_n(k)$ in location $k$ of block $DSSD \cdots D$, and so on;
- the tree of squares, which contains $\sum_{n=0}^{N-1} [SDSD \cdots Dx_n(k)]^2$ in location $k$ of block $SDSD \cdots D$, and so on.

Figure 4: Example graph basis: the Haar wavelets
Computing all the coefficients of all the blocks in an $L$-level tree starting from $d$ samples takes $O(dL) = O(d \log d)$ operations per random vector, for a total of $O(Nd \log d)$ operations. After we do this for all the random vectors $X_n$ in the ensemble $X$, we can produce the binary tree of variances by using Equation 2: at index $k$ of block $DSD \ldots S$, for example, it contains
\[
\sigma_{DSD \ldots S}^2(k) = \frac{1}{N} \sum_{n=0}^{N-1} DS \ldots DX_n(k)^2 - \left( \frac{1}{N} \sum_{n=0}^{N-1} DS \ldots DX_n(k) \right)^2
\]
This is the variance of the wavelet packet coefficient defined by the filter sequence $DSD \ldots S$ and the location $k$. Forming this tree takes an additional $O(d \log d)$ operations. For convenience we introduce the following notation for these blocks which compactly encodes filter sequences such as $DSDSSD$. Let $n = (n_{L-1}, \ldots, n_1, n_0)$ be the binary representation of the nonnegative integer $0 \leq n < 2^L$. Then $n = n_{L-1} 2^{L-1} + \cdots + n_1 2^1 + n_0$ with $n_i \in \{0, 1\}$ for $i = 0, 1, \ldots, L - 1$. This will represent the filter sequence $F_0 F_1 \ldots F_{L-1}$, where
\[
F_i = \begin{cases} S, & \text{if } n_i = 0, \\ D, & \text{if } n_i = 1. \end{cases}
\]
Block $n$ at level $j$ of the variance tree will be denoted $V_j n$; its coefficients form the sequence which may be denoted
\[
(\sigma_j n(X)(0), \sigma_j n(X)(1), \ldots, \sigma_j n(X)(d/2^j - 1)).
\]
Notice that the two children of block $V_j n$ are the $S$ block $V_{j+1} n$ and the $D$ block $V_{j+1} n + 1$, both at level $j + 1$.

The tree of variances may now be searched for the graph basis which minimizes $H$. We begin by defining the information cost function:
\[
\mathcal{H}(V_j n) = \sum_{k=0}^{d/2^j - 1} \log \sigma_j n(X)(k).
\]
Since this calculation will always be done at some fixed finite precision $\epsilon > 0$, we never really have a singularity and we will simply replace $\log 0$ by $\log \epsilon$.

Now let $I_j n$ be the information cost of the best graph basis in the subtree whose root is $V_j n$. Let $U_j n$ be the collection of blocks in that best graph basis. In other words, $\sum_{V \in U_j n} \mathcal{H}(V)$ is minimal with all the blocks $V = V_j n'$ in the graph $U_j n$ being descendants of $V_j n$ and satisfying $j' \geq j$. Then we can find the minimum for the whole tree by the following preorder recursion:

- Compute $I_j n = \mathcal{H}(V_j n)$.
- If $j = L$, then set $U_L n = \{ V_L n \}$ and return.
- Find $U_{j+1} n$ and $I_{j+1} n$ recursively.
- Find $U_{j+1} n + 1$ and $I_{j+1} n + 1$ recursively.
- If $I_j n \leq I_{j+1} n + I_{j+1} n + 1$ then $V_j n$ by itself gives as good a representation of $X$ as $U_{j+1} n \cup U_{j+1} n + 1$, so
  - Retain $I_j n$ as the information cost;
  - Set $U_j n = \{ V_j n \}$, a single-block graph.
- Else $I_j n > I_{j+1} n + I_{j+1} n + 1$, so it is cheaper to represent $X$ using the descendant blocks in $U_{j+1} n \cup U_{j+1} n + 1$, so
  - Set $I_j n = I_{j+1} n + I_{j+1} n + 1$;
  - Set $U_j n = U_{j+1} n \cup U_{j+1} n + 1$, the union of the best graph bases for the descendants of $V_j n$.
- Return.
This recursion terminates leaving the best graph basis in \( U_{00} \) and the lowest information cost of any graph basis in \( I_0 \). The rather straightforward proof of these facts can be found in [3].

Notice that each block is examined twice: once when it is a parent and once when it is a child. This means that the best-basis search requires as many comparison operations as there are blocks in the tree, which is \( O(d) \). Computing \( \mathcal{H} \) costs no more than a fixed number of arithmetic operations per coefficient in the tree, which is \( O(d \log d) \). Thus the total cost of the search is \( O(d \log d) \).

\( U_{00} \) will be called the **joint best basis** for the ensemble \( X \), in the Haar–Walsh wavelet packet library. We can denote by \( U \) the \( d \times d \) orthogonal matrix which corresponds to the orthonormal basis \( U_{00} \). Abusing notation, we write \( \{ U_i \in \mathbb{R}^d : i = 1, \ldots, d \} \) for the rows of \( U \). We may suppose that these rows are numbered so that \( \sigma(UX) \) is in decreasing order; this can be done by sorting all the \( d \) coefficients in all the blocks \( V \in U_{00} \) into decreasing order, which can be done in \( O(d \log d) \) operations.

If we fix \( \epsilon > 0 \) and let \( d' \) be the smallest integer such that

\[
\sum_{n=1}^{d'} \sigma(UX)(n) \geq (1 - \epsilon) \text{Var}(X),
\]

then the projection of \( X \) onto the row span of \( U' = \{ U_1 \ldots U_{d'} \} \) contains \( 1 - \epsilon \) of the total variance of the ensemble \( X \). Call this projected ensemble \( \tilde{X} \). The \( d' \) row vectors of \( U' \) are already a good basis for the ensemble \( \tilde{X} \), but they may be further decorrelated by Karhunen–Loève factor analysis. The row vectors of \( U' \) are just \( U'_i = U_i \) for \( 1 \leq i \leq d' \), and the autovariance matrix for this new collection is given by

\[
M'_{ij} = \frac{1}{N} \sum_{n=1}^{N} U'_i \tilde{X}_n U'_j \tilde{X}_n.
\]

Here \( \tilde{X}_n = \tilde{X}_n - E(\tilde{X}) \) is a vector in \( \mathbb{R}^d \), and \( E(\tilde{X}) = 0 \). Thus \( M' \) is a \( d' \times d' \) matrix and can be diagonalized in \( O(d'^3) \) operations. Let \( K' \) be the matrix of singular vectors of \( M' \). Then \( K'' \) changes from the joint best basis coordinates (calculated from the standard coordinates by \( U' \)) into coordinates with respect to these decorrelated singular vectors. We may thus call the composition \( K''U' \) (associated to \( \epsilon \)) the approximate Karhunen–Loève transform with relative variance error \( \epsilon \).

### 3.1 Complexity

The algorithm is fast because we expect that even for small \( \epsilon \) we will obtain \( d' \ll d \). To count operations in the worst case, assume:

1. \( N \) random vectors;
2. \( d \)-dimensional parameter space \( \mathbb{R}^d \);
3. Full-rank autovariance matrix (\( N > d \)).

**Finding the approximate Karhunen–Loève basis**

- Expanding \( N \) vectors \( \{ X_n \in \mathbb{R}^d : n = 1, 2, \ldots, N \} \) into wavelet packet coefficients: \( O(Nd \log d) \);
- Summing squares into the variance tree: \( O(d \log d) \);
- Searching the variance tree for a best basis: \( O(d + d \log d) \);
- Sorting the best basis vectors into decreasing order: \( O(d \log d) \);
- Diagonalizing the autovariance matrix of the top \( d' \) best-basis vectors: \( O(d'^3) \).
Adding these up, we see that the total complexity of constructing the approximate Karhunen-Loève transform \( K''U' \), is \( O(Nd \log d + d^3) \). This compares favorably with the complexity \( O(Nd^2 + d^3) \) of the full Karhunen-Loève expansion, since we expect \( d' \ll d \).

Depending on circumstances, the last step \( U' \mapsto K''U' \) may not be necessary, since a large reduction in the number of parameters is already achieved by transforming into the orthonormal basis determined by \( U' \). This reduces the complexity to \( O(Nd \log d) \), with the penalty being less decorrelation of the factors.

The approximate Karhunen-Loève transform of 1 vector

- Computing the wavelet packet coefficients of 1 vector: \( O(d \log d) \).
- Applying the \( d' \times d' \) matrix \( K'' \): \( O(d^2) \).

Since \( d' \ll d \), this estimate compares favorably with the complexity of applying the full Karhunen-Loève transform to a vector, which is \( O(d^2) \). Further savings are possible, notably because only a small fraction \( d'' \ll d' \) of the Karhunen-Loève singular vectors are needed to capture almost all of the original ensemble variance. Hence we can take \( K'' \) to be just the first \( d'' \) of the columns of \( K' \), and then the total complexity of applying \( K''U' \) is bounded by \( O(d \log d + d''^2) \).

If we expect to update the Karhunen-Loève basis, then we might also expect to update the average vector and the average value of each coordinate in the library of bases, as well as the variance of the ensemble. But since we keep a sum-of-squares tree and a sum-of-coefficients or means tree rather than a variance tree, each additional random vector just contributes its wavelet packet coordinates into the means tree and the squares of its coordinates into the sum-of-squares tree. The variance tree is updated at the end using the correct new means. This results in the following update complexity:

Updating the approximate Karhunen-Loève basis

- Expanding 1 vector into wavelet packet coefficients: \( O(d \log d) \).
- Adding the coefficients into the means tree: \( O(d \log d) \).
- Adding the squared coefficients into the squares tree: \( O(d \log d) \).
- Forming the variance tree and computing the new information costs: \( O(d \log d) \).
- Searching the variance tree for the joint best basis: \( O(d + d \log d) \).

So one new vector costs \( O(d \log d) \), and updating the basis with \( N > 1 \) new vectors costs \( O(Nd \log d) \).

4 Classification in large data sets

The Karhunen-Loève transform can be used to reparameterize a problem so as to extract prominent features with the fewest measurements. When the number of measurements is huge, the fast approximate algorithm must be used at least as a “front end” to reduce the complexity of the SVD portion of finding the Karhunen-Loève basis.

We list a few examples to give some indication of the size of problem that can be treated by the approximate method on typical tabletop computing equipment.
4.1 The rogues’ gallery problem

The “rogues’ gallery” problem is to identify a face from among a collection of faces. This problem was first suggested to me by Lawrence Sirovich, who also provided the data used in this experiment. The random vectors were several thousand digitized 128x128x8-bit pictures of Brown University students, so \( d = 128^2 = 16,384 \). These were initially normalized with the pupils impaled on two fixed points near the center of the image. In [6, 5], a supercomputer was used to compute the Karhunen-Loève transform either of the complete set of pixels or else of an oval subset centered about the eyes. In the following, we will follow Sirovich’s methodology and nomenclature, only we will replace the Karhunen-Loève transform with the lower-complexity approximate algorithm.

![Face images](image)

Figure 5: Face, minus the average face, yields a caricature.

For the experiment described below, we start with a more limited data set containing 143 pictures. Since the ensemble was fixed, we could subtract the average vector at the outset. Thus we transformed the data to floating point numbers, computed average values for the pixels, and then subtracted the average from each pixel to obtain “caricatures,” or deviations from the average. Figure 5 is one of these caricatures.

The left graph in Figure 6 shows how the variance accumulates pixel by pixel, with the pixels sorted into decreasing order of variance.

![Graph](image)

Figure 6: Accumulation of variance in the original basis and the joint best basis.

Each caricature was treated as a picture and expanded into 2 dimensional wavelet packets as described in [7]. The squares of the amplitudes were summed into a tree of variances, which was then searched for the joint
best basis. In the joint best basis, 400 coordinates (of 16,384) contained more than 90% of the variance of the ensemble.

The right graph in Figure 6 shows the accumulation of total variance on the first $d'$ coordinates in the joint best basis, sorted in decreasing order, as a fraction of the total variance of the ensemble, for $1 \leq d' \leq 1000$. Using 1000 parameters captures more than 95% of the ensemble variance, but requires somewhat more computer power than is readily available on a desktop. A 400 parameter system, on the other hand, can be analyzed on a typical workstation in minutes so we choose $d' = 400$.

The top 400 coordinates were recomputed for each caricature and their autocovariance matrix over the ensemble was diagonalized using the LINPACK singular value decomposition routine.

Figure 7 shows the accumulation of total variance on the first $d''$ coordinates in the approximate Karhunen–Loève basis, sorted in decreasing order, as a fraction of the total variance of the 400 joint best basis coefficients, for $1 \leq d'' \leq 143$. The Karhunen–Loève post-processing for this small ensemble concentrates 98% of the retained variance from the top 400 joint best-basis parameters into 10 coefficients.

### 4.2 The fingerprint classification problem

Virtually the same method as the one applied to faces can be applied to fingerprints for identification purposes. The United States FBI uses 8 bits per pixel to define the shade of gray and stores 500 pixels per inch, which works out to about 900,000 pixels and 0.6 megabytes per finger to store fingerprints in electronic form. This means that $d \approx 10^6$, so we are forced to use the fast approximate algorithm if we wish to compute the Karhunen–Loève expansion of a fingerprint.

There is no apparent relation between the parameters chosen by the Karhunen–Loève transform and the traditional parameters (location of “minutiae” points in a fingerprint) which are used by police inspectors to describe fingerprints. Thus the additional classifying values would have to be stored alongside the more traditional values. No one is likely to complain, though, since the Karhunen–Loève parameters occupy only a few hundred bytes, which is a negligible amount of space compared to the million bytes of raw data.

### 4.3 Rank reduction for complex classifiers

The principle behind the fast approximate Karhunen–Loève transform is to employ a relatively low-complexity $O(d^2 \log d)$ “front end” to reduce the rank of the subsequent high-complexity $O(d^3)$ algorithm from $d$ to $d' \ll d$. 

![Figure 7: Accumulation of variance in the approximate Karhunen–Loève basis.](image-url)
If some measurements on an ensemble of random vectors are to be processed further by some other complex algorithm, we may similarly gain a significant speed advantage by pre-processing the data to reduce the number of parameters. A typical example would be processing for statistical classification from a large set of measurements. Classes, or regions in the measurement space \( \mathbb{R}^d \), may have complicated boundaries which must be approximated by high-order polynomial hypersurfaces. Deciding at high orders whether a point lies in a particular region becomes very expensive when the dimension \( d \) grows large, so a reduction in the number of parameters will gain speed even if does not by itself simplify the geometry of the regions.

High complexity classifiers are used in speech recognition and machine vision systems. Adding a front end to reduce their workload is like adding a hearing aid or glasses so they can better focus on the most-varying features. In some cases, the speed is desirable because we wish to perform the classification in “real time” or least fast enough to keep up with the inflowing data. Some examples are:

- Mechanical failure detection from strain gauge data: \( d \approx 10^2 \);
- Target recognition from 1-dimensional radar range profiles: \( d \approx 10^3 \);
- Detection of irregular heartbeats from acoustic samples: \( d \approx 10^3 \);
- Phoneme detection: \( d \approx 10^3 \);
- Optical character recognition: \( d \approx 10^3 \);
- Detection of machine tool breakage from acoustic samples: \( d \approx 10^3 \);

5 Jacobsians of complicated maps

Suppose that \( T : \mathbb{R}^p \to \mathbb{R}^d \) is some smooth vector field with \( p \ll d \). We may think of making plenty \( (d) \) of measurements of \( Tx \) for a variable \( x \) with only a few \( (p) \) degrees of freedom. This situation models one course of action to determine what a complicated map \( T \) is doing.

5.1 Approximating the tangent space in principal components

Recall that the Jacobian of \( T \) at a point \( x \in \mathbb{R}^p \) is the \( d \times p \) matrix \( J = J_T[x] \) which gives the best linear approximation to \( T \) in a neighborhood of \( x \). The coefficients of \( J \) are the various partial derivatives of \( T \).

\[
J_T[x](i,j) \overset{\text{def}}{=} \lim_{r \to 0} \left\{ e_i, \frac{T(x + re_j) - T(x)}{r} \right\},
\]

Here \( 1 \leq i \leq d, 1 \leq j \leq p \), and \( e_j \) is the \( j \)th standard basis vector. However, the numerical computation of this Jacobian poses some difficulties because the difference quotient is ill-conditioned. Furthermore, the Jacobian might itself be an ill-conditioned matrix, but this difference quotient procedure offers no way of estimating the condition number of \( J \). We will address these difficulties by replacing the difference quotient formula with an approximation based on the Karhunen–Loève transform for the positive matrix \( J^*J \). The error will lie solely in the approximation, since the Karhunen–Loève transform is orthogonal and thus perfectly conditioned. We will estimate the condition number of \( J \) from the singular value decomposition of \( J^*J \). Then

\[
\text{cond}(J) = \sqrt{\text{cond}(J^*J)} \approx \sqrt{\mu_1/\mu_p},
\]

where \( \mu_1 \) and \( \mu_p \) are respectively the first and \( n \)th singular values of our estimate for \( J^*J \).

Suppose first that \( T \) is a linear map, so that \( T = J \) is its own Jacobian. Fix \( x \in \mathbb{R}^p \), and consider the ball \( B_r = B_r(x) \overset{\text{def}}{=} \{ y \in \mathbb{R}^p : \| y - x \| \leq r \} \) of radius \( r > 0 \), centered at \( x \). We can consider the image
$J B_r = \{ J y : y \in B_r(x) \} \subset \mathbb{R}^d$ of this ball under the transformation $J$ to be an ensemble of random vectors. This will have expectation $E(J B_r) = J E(B_r) = J x$, and we can compute the autocovariance matrix of the zero-mean ensemble $\mathcal{J}B_r \overset{\text{def}}{=} J B_r - J x$ as follows:

$$E(\mathcal{J}B_r \otimes \mathcal{J}B_r) = E_{y \in B_r(x)} \langle J y | J y^* \rangle = J E_{y \in B_r(x)} \langle y | y^* \rangle J^* = r^2 J J^*.$$

Here $y = x - x$ and the last equality holds since $E_{y \in B_r(x)} \langle y y^* \rangle = r^2 I_d$ is just a constant times the $d \times d$ identity matrix and thus commutes out from between $J$ and $J^*$. Thus $r^{-2} E(\mathcal{J}B_r \otimes \mathcal{J}B_r) = J J^*$.

**Proposition 1** For every matrix $J$,

$$\text{Rank} \; J J^* = \text{Rank} \; J = \text{Rank} \; J^* = \text{Rank} \; J J^*.$$

**Proof:** To prove the first equality, notice that Range $J J^* \subset$ Range $J$ implies that Rank $J \geq$ Rank $J J^*$. Now suppose that Range $J J^* \neq$ Range $J$. There is some $y \neq 0$ with $y \in$ Range $J$ and $\langle y, J J^* z \rangle = \langle J^* y, J^* z \rangle = 0$ for all $z$. Putting $z = y$ we see that $J^* y = 0$. But by its definition, $y = J x$ for some $x$, so we have $\| y \|^2 = \langle y, x \rangle = \langle J^* y, x \rangle = 0$, a contradiction. The third equality follows from the same argument if we substitute $J^*$ for $J$. For the middle equality, note that Rank $AB \leq \min(\text{Rank} \; A, \text{Rank} \; B)$, so that the first equality gives Rank $J \leq$ Rank $J^*$ while the third gives Rank $J \geq$ Rank $J^*$. \hfill \Box

Suppose $J$ is an $d \times p$ matrix with $d \geq p$. If $J$ has maximal rank $p$, then $J^* J$ also has rank $p$. Now, the condition number of $J$ is

$$\left(\sup_{x \neq 0} \frac{\| J x \|^2}{\| x \|^2} \right) \left(\inf_{y \neq 0} \frac{\| J y \|}{\| y \|}\right)^{-1}. \tag{15}$$

If $J^* J$ has $n$ nonzero singular values $\mu_1 \geq \cdots \geq \mu_p > 0$, counting multiplicities, then the supremum is $\sqrt{\mu_1}$ and the infimum is $\sqrt{\mu_p}$. To see this, let $z_1, \ldots, z_p$ be the orthonormal basis of $\mathbb{R}^p$ consisting of unit singular vectors for $J^* J$, which is guaranteed to exist because $J^* J$ is a Hermitian matrix. Writing $x = \sum_i a_i z_i$ we have

$$\frac{\| J x \|^2}{\| x \|^2} = \frac{\langle J x, J x \rangle}{\| x \|^2} = \frac{\langle J^* J x, x \rangle}{\| x \|^2} = \frac{\sum_i a_i^2 \mu_i}{\| x \|^2}. \tag{16}$$

This average is maximized when just $a_1$ is nonzero and minimized when just $a_p$ is nonzero; it then equals $\mu_1$ and $\mu_p$, respectively. Thus we can compute the condition number of $J$ using the formula in Equation 14.

Now suppose that $T$ is any smooth vector field from $\mathbb{R}^p$ to $\mathbb{R}^d$, and $x$ is some point in $\mathbb{R}^p$. We compute $z_r = E(T B_r)$, where $T B_r = T B_r(x) = \{ y : \| y - x \| \leq r \}$; this is the expected value of $T y$ for $y$ in the ball $B_r(x)$ of radius $r$ centered at $x$. This average gives a second-order approximation to $T x$.

**Proposition 2** \(\| z_r - T x \| = O(r^2)\) as \(r \to 0\).

**Proof:** We can use Taylor's theorem to write $T(x + y) = T x + J y + O(\| y \|^2)$ for $\| y \| \leq r$. But $E(J y) = J E(y) = 0$ because we evaluate the expectation over $y \in B_r(0)$. Thus $E(TB_r) = T x + O(r^2)$.

We now define a positive matrix

$$A = A_r = E([TB_r - z_r] \otimes [TB_r - z_r]), \tag{17}$$

where the expectation is taken over the ball of radius $r$. Our main theorem is the following:

**Theorem 3** \(\lim_{r \to 0} \frac{1}{r^2} A_r = J J^*\).
Figure 8: Tangent space (Range $J$) of $T$ and its approximation (Range $A_r$)

Proof: Using Proposition 2 we write $z_r = T x + O(r^2)$. We then use Taylor’s theorem to get the following estimate:

$$[T(x+y) - z] \otimes [T(x+y) - z] = \left[ J y + O(\|y\|^2) \right] \otimes \left[ J y + O(\|y\|^2) \right] = (J y)(J y)^* + O(\|y\|^3).$$

Taking the expectation of both sides over $y \in B_r(0)$ gives $A_r = r^2 J J^* + O(r^3)$, yielding the desired result. Notice that we also get an error estimate: $\|J J^* - \frac{1}{r^2} A_r\| = O(r)$. □

Now suppose that $x \in \mathbb{R}^p$ is a point for which the Jacobian $J_T[x]$ has full rank $p$. Full rank is an open condition, i.e., it is possessed by all matrices sufficiently close to $J$ as well, so we have the following reassuring fact:

**Corollary 4** For all sufficiently small $r > 0$, Rank $A_r \geq$ Rank $J = p$. □

Then we can approximate the map $T$ in a neighborhood of $T x$ using the singular vectors for $A_r$:

**Corollary 5** If $\{z_1, \ldots, z_p\} \subset \mathbb{R}^d$ is a set of unit orthogonal singular vectors for $A_r$, then there are $p$ linear functions $c_1, \ldots, c_p$ on $\mathbb{R}^p$ such that $T(x + y) = z + \sum_{i=1}^p c_i(y) z_i + O(r\|y\|^2)$. □

We are not really concerned with the rank of $A_r$ being too small, since $A_r$ is a $d \times d$ matrix and $d \gg p$ is the interesting situation. Rather, we worry that choosing a too-large value for $r$ will result in Rank $A_r$ being too large, so that we will not be able to identify the top few singular vectors which approximately span Range $J$. The problem is that if $T$ has nonvanishing higher-order derivatives, then the range of $A_r$ will be higher-dimensional than the tangent space to $T x$ at $T x$, which is the range of $J$. Schematically, this is depicted in Figure 8. The range of $A_r$ is drawn as the two unit singular vectors $z_1, z_2$ multiplied by their respective singular values $\mu_1, \mu_2$. Notice that $\mu_2 \ll \mu_1$, illustrating what happens with smooth $T$: the variation $\mu_2$ of $TB_r$ in the nontangential direction $z_2$ is much smaller than the variation $\mu_1$ in the tangent or $z_1$ direction. In practice, we will always have Rank $A_r = d$ because to a finite precision machine every approximate matrix looks like it has full rank. However, if we arrange the singular values of $A_r$ (with multiplicity) in decreasing order $\mu_1 \geq \cdots \geq \mu_p \geq \cdots \geq 0$, then for small enough $r$ we expect a steep drop between $\mu_p$ and $\mu_{p+1}$. This then provides a method of choosing the largest $r$ for which the singular vectors of $A_r$ provide an accurate parametrization of $T$ near $x$. Namely, we let $r$ increase until
\sqrt{\mu_{p+1}/\mu_p} \text{ reaches some preset threshold of precision } \epsilon_p. \text{ Then the nontangential components will contribute an error which is } 1/\epsilon_p \text{ times smaller than the tangential components.}

The functions \( c_1, \ldots, c_p \) in Corollary 5 correspond to partial derivatives, but they are computed by orthogonal projection. We define matrix coefficients \( c_{ij} \) using the elementary basis vectors \( e_j \) of \( \mathbb{R}^p \) as follows:

\[
C_{ij} \overset{\text{def}}{=} \frac{1}{r} (\langle z_i, T(x + re_j) \rangle - z) \tag{18}
\]

Then we extend this to the superposition \( y = \sum_{j=1}^{p} a_j e_j \) by taking linear combinations as follows:

\[
c_k(y) \overset{\text{def}}{=} \sum_{j=1}^{p} C_{kj} a_j. \tag{19}
\]

Comparing Equation 18 with Equation 13, we see that the \( d \times p \) matrix \( J \) has been replaced with the \( p \times p \) matrix \( C \), the limit has been reduced to a single evaluation using the largest acceptable \( r \), the initial point \( T_x \) is now an average \( z \), and the standard basis \( \{e_j : j = 1, \ldots, d\} \) in the range space has been replaced with a new orthonormal basis which is locally adapted to \( T \).

Notice that the columns of the \( p \times p \) matrix \( C = (C_{ij}) \) are given by the top \( p \) coordinates of the Karhunen–Loève transform of the secant vectors

\[
\frac{1}{r} [T(x + re_1) - z], \frac{1}{r} [T(x + re_2) - z], \ldots, \frac{1}{r} [T(x + re_p) - z],
\]

since the unit singular vectors \( z_1, \ldots, z_p \) of \( A_r \) are the Karhunen–Loève eigenvectors for the ensemble \( TB_r \). These secant vectors are approximations to the directional derivatives of \( T \) in the directions \( e_1, e_2, \ldots, e_p \), and the Karhunen–Loève transform projects them onto the principal orthogonal components along Range \( T \).

5.2 Fast approximate Jacobians

We can use the fast approximate Karhunen–Loève algorithm to compute the approximate Jacobian. Suppose that the domain of \( T \) includes a cube centered at the origin, namely \( B_r = B_r(0) \overset{\text{def}}{=} \{x \in \mathbb{R}^d : |x_1| \leq r, \ldots, |x_p| \leq r\} \). Suppose we lay down a uniform grid of points of the form \( x_k = k \) where \( k = 0, \pm 1, \pm 2, \ldots \) and \( i = 1, \ldots, p \). The intersection of the cube with this grid, which we will also call \( B_r \), contains \( (2r+1)^p \) points in all. We now compute \( T_x \) at all points \( x \in B_r \), and call the resulting set \( TB_r \). This will be our ensemble of “random” vectors. Each \( x \in B_r \) produces a vector \( T x \in \mathbb{R}^d \) which requires \( d \) numbers to store, so \( TB_r \) will contain \( |B_r| d = (2r+1)^p \times d \) floating-point numbers.

The approximation to \( T \) at 0 using \( B_r \) will be computed by the wavelet packet best-basis algorithm above. The mean vector \( z = E(TB_r) \) is computed in all wavelet packet bases at once in the means tree. We form the variance tree from the squares tree and the means tree, and search it for the joint best basis of \( TB_r \). We may assume that this basis is sorted into decreasing order of variance.

Now we take \( d' \) of the most-varying terms out of the \( d \) in the joint best basis to retain \( 1 - \epsilon \) of the total variance. We then form the \( d' \times d' \) autocovariance matrix for these \( d' \) joint best basis vectors and diagonalize it, finding the singular vectors \( z_1, \ldots, z_{d'} \) and corresponding singular values \( \mu_1 \geq \ldots \geq \mu_{d'} \). We put the singular vectors into the columns of a matrix \( K' \), and get the approximate Karhunen–Loève transform \( K'' \), a \( d' \times d' \) matrix which gives the approximate principal components when applied to the top \( d' \) joint best basis coordinates.

We now test whether the cube \( B_r \) is too large for the singular vectors to well-approximate the tangent space. The rank of the Jacobian is at most \( p \), so we must have \( \epsilon_p(r) = \sqrt{\mu_{p+1}/\mu_p} \ll 1 \). This gives the first parameter of the algorithm: we know that \( \epsilon_p(r) \to 0 \) as \( r \to 0 \), so if it exceeds a preset threshold we need to reduce \( r \). However, if \( \epsilon_p(r) \) meets our requirements, then we can discard all but the first \( p \) columns of \( K' \) to get the \( d' \times p \)
matrix $K''$. The range of $K''$ serves as the approximate tangent space at $T_0$, and $K'''$ computes coordinates in this space from the top $d'$ joint best basis coordinates.

Finally, we form the approximate Jacobian into this approximate tangent space by using Equation 18 with an approximate principal factor for $z$. One by one, we take the secant vectors $\frac{1}{2} |T(x + r e_j) - z|$ for $j = 1, 2, \ldots, p$, which live in $\mathbb{R}^d$, and we find their joint best basis expansions. We then extract from each of those expansions the previously-chosen $d'$ coordinates which have most of the variance over the ensemble and apply $K'''$ to the vectors of these coordinates. That gives a list of $p$ vectors in $\mathbb{R}^p$ which approximate the coefficients of the partial derivatives $\partial_1 T, \ldots, \partial_p T$ in the approximate tangent space basis.

The approximate Jacobian data consists of

- the joint best-basis description down to $d'$ coordinates ($d'$ numbers),
- the $p$ vectors of the approximate tangent space expressed as combinations of the first $d'$ joint best basis vectors ($pd'$ numbers), and
- the $p \times p$ matrix $C$ of partial derivatives expressed as combinations of the approximate tangent vectors.

Computing these quantities will cost us $O(|B_r| \times [d'^3 + d'^2 \log d'])$ arithmetic operations, where we expect $d' \ll d$.

### 5.3 Efficient storage of complicated maps

Suppose for this application that the domain of $T$ is the unit cube in $Q \subset \mathbb{R}^p$ defined by $Q = \{ x \in \mathbb{R}^p : 0 \leq x_1 \leq 1, \ldots, 0 \leq x_p \leq 1 \}$, and suppose we lay down a uniform grid of points of the form $x_i = r/R$ where $r = 0, 1, \ldots, R$ and $i = 1, \ldots, p$. We will call this grid $G$; it has mesh $1/R$ and contains $|G| = (R+1)^p$ points in all. We now compute $T x$ at all points $x \in G$, and call the resulting set $TG$. This is an enormous data set: each $x \in G$ produces a vector $T x \in \mathbb{R}^d$ which requires $d$ numbers to store, so $TG$ will contain $|G| d = (R+1)^p d$ floating-point numbers.

We now use approximate Jacobians to reduce the size of the data set. We will do this by building up patches in the domain where $T$ is well-approximated by its approximate Jacobian. With the fast update algorithm, we
can segment the domain of $T$ into patches on which we are sure that the approximation remains within a preset error. Suppose we start at $0 \in G$, at one corner of the cube $Q$. For each $r = 0, 1, 2, \ldots$ we define the set $Q_r = Q_r(0) \overset{\text{def}}{=} \{ x \in G : 0 \leq x_i \leq r/R \text{ for } i = 1, \ldots, p \}$. We also define the set $P_r = P_r(0) \overset{\text{def}}{=} Q_r(0) \setminus Q_{r-1}(0)$, the partial cutical shell at radius $r/R$. Then $Q_{r+1} = Q_r \cup P_{r+1}$. This arrangement is depicted for the two-dimensional case $p = 2$ by the points enclosed in lightly-colored boxes in Figure 10. Note that $Q_r$ contains $|Q_r| = (r + 1)^p$ points, while $P_r$ contains $|P_r| = |Q_r| - |Q_{r-1}| = (r + 1)^p - r^p \approx p r^{p-1}$ points.

The segmentation algorithm works by first initializing $r = 1$, and then iterating through an algorithm that enlarges a patch on which $T$ is linearly approximated by its approximate Jacobian. We stop enlarging the patch $Q_r$ when the variance of $TQ_r$ along the approximate tangent vectors to $T$ stops being much larger than the variance along approximate normal vectors. In following, we assume that the approximate Karhunen–Loève basis for $TQ_{r-1}$ has already been determined, using the algorithm in Section 3 above. The update algorithm is also defined in that section.

- Compute the wavelet packet means and sums-of-squares trees for the extra vectors $TP_r$;
- Update the joint best basis for $TQ_{r-1}$ by adding in the data from $TP_r$ to get the joint best basis for $TQ_r$;
- Compute and store the approximate Karhunen–Loève basis $K_r^*$ for $TQ_r$ and the singular values $\mu_1 \geq \ldots \geq \mu_{p^r} \geq 0$;
- If $\epsilon_\mu(r) = \sqrt{\mu_{p^r+1}/\mu_p}$ is too large, then:
  - Compute and store $z = ETQ_{r-1}$ for the center;
  - Compute and store the approximate tangent vectors $K_{r-1}^*$ for the patch $TQ_{r-1}$;
  - Compute and store the approximate Jacobian using $K_{r-1}^*$ and Equation 18;
  - Reset $k = 1$;
  - Move to the next free point in $G$;
- Else if $\epsilon_\mu(r)$ is still small enough, then increment $r$ by 1;
- Repeat.

This algorithm will eat away at the domain $G$, producing a covering of patches $Q$ of various sizes, each with its center point $z_Q = ETQ$, its local approximate tangent space $K_Q^o$ and its approximate Jacobian $C_Q$. These quantities will require $d'$, $pd'$ and $p^2$ real numbers to store, respectively. If there are a total of $N$ patches, then the total amount of data to store is $N(d' + pd' + p^2) = O(Npd')$ numbers since $p \leq d'$. If $p$ is small, $N \ll |G|$ and $d' \ll d$, then this compares favorably with the storage requirements for $TG$, namely $O(d|G|)$.

The complexity of computing these quantities on all of the patches can be estimated from the complexity of finding the approximate Jacobian for the a single patch containing all of $G$, since this if the worst case. But from the previous section, we see that this is $O(|G| \times [d'^3 + d^2 \log d])$ arithmetic operations.

Applying this approximation of $T$ to a vector $x \in \mathbb{R}^p$ involves first finding the patch $Q$ with $x \in Q$. Suppose that $x_Q$ is the center grid point of $Q$. Then in the first $d'$ joint best-basis coordinates,

$$\tilde{T}x = z_Q + K_Q^o C_Q (x - x_Q).$$

We finally superpose the $d'$ joint best basis vectors to get the coordinates of the point $Tx \in \mathbb{R}^d$ from $\tilde{T}x$. The complexity of computing $Tx$ this way is $O(p + p^2 + pd' + d' + d \log d)$ which under our assumptions is bounded by $O(d \log d)$.

### 5.4 Precomputation for inverse problems

The final application is to use the local approximate Jacobians to invert the map $x \mapsto Tx$, which we suppose has already been computed at all points $x$ on a finite grid $G$.

One classical way is to use linear interpolation: given $y \in \mathbb{R}^d$, we find the nearest points $Tx_k \in \mathbb{R}^d$ computed from grid points $x_k \in G$ and write $y = \sum_k a_k Tx_k$. Then the linear approximation to $T^{-1} y$ is just $\sum_k a_k x_k$. This is exactly correct for linear maps $T$ and has at least $O(h)$ accuracy for a differentiable map $T$ on a grid with mesh $h$. However, it requires that we store the precomputed values $\{Tx : x \in G\}$ and that we search the whole list for the points close to $y$. The last step, in particular, requires computing $|G|$ distances for a grid $G$.

If we have invested the effort to compute the approximate Jacobian representation of $T$, then the inverse can be approximated from that data instead. Let $N$ be the number of patches in the cover of $G$, and suppose that $N \ll |G|$. We also suppose for the sake of simplicity that we keep the same number $d'$ of joint best basis components in each patch, although of course they may be different components in different patches. Finally, we suppose that we have already computed the inverses $C_Q^{-1}$ of the approximate Jacobians on each patch $Q$, which requires a one-time investment of $O(Np^3)$. Then the necessary computations and their complexities for computing $T^{-1} y$ at a single $y \in \mathbb{R}^d$ are the following:

- Find the complete wavelet packet expansion of $y$, which simultaneously computes all joint best-basis expansions $\tilde{y} : O(d \log d)$;
- Compute the distances from $\tilde{y}$ to the means $z_Q$ for each patch $Q$, and let $Q$ henceforth be the patch with nearest mean: $O(Nd')$;
- Compute the approximate inverse,

$$T^{-1} y \approx x_Q + C_Q^{-1} K_{Q'} (\tilde{y} - z_Q)$$

for the nearest-mean patch $Q$: $O(d' + pd' + p^2 + p) = O(pd')$. 

References


